



A tale of two transforms

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Aims

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- Briefly to compare Khmaladze transforms 1 and 2
- to give some intuition underlying KT2
- to apply KT2 to simple simulated Bernoulli trials:
- to compare practicability of KT1 and KT2 in testing gof in distn free manner

Both KT's purport to test gof by distn free methods

- It is the best of times
 - Elegant and powerful transforms of stochastic processes are available
 - ostensibly very useful methodologies

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 - Elegant and powerful transforms of stochastic processes are available
 - ostensibly very useful methodologies
- It is the worst of times
 - Almost no-one is applying these transforms

KT1

- The basic empirical process

$$v_n(t) = \frac{1}{\sqrt{n}} [F_n(x) - F_\theta(x)]$$

is in the nature of a Brownian bridge.

- Usual projection is not invertible

$$u_n(x) = v_n(x) - g(x, \theta) \int h(y, \theta) dv_n(y) + o_P(1)$$

- But this transform from BB to Brownian motion is invertible

$$dv(t) = -\frac{v(t)}{1-t} dt + dw(t)$$

$$v(t) = (1-t) \int_0^t \frac{1}{1-s} dw(s)$$

This gives rise to the KT1

$$dw_{n,\theta}(x) = \sqrt{n} \left[dF_n(x) - h_\theta(x) dF_\theta(x) \Gamma_{x,\theta}^{-1} \int_{y=x}^{\infty} h_\theta(y) dF_n(y) \right]$$

This is a regression of $dF_n(x)$ on $\int_{y=x}^{\infty} h_\theta(y) dF_n(y)$

The limiting distn of $w_{n,\theta}(x)$ is indeed a BM, so we have achieved a distn free test. But implementation is awkward:

- hard to program: two interlocking integrals
 - two inter-related integrals, one over 'future' values
 - coding is very dependent on distributional family chosen
- underlying maths is technical, and not so intuitive
- cannot be simulated in real time; assume ML estimates known from whole sample
- modern tendency to eschew theory in favour of numerical methods, applied blindly through statistical packages

D = data, F = data generating process;
v = empirical process; w = BM

KT1: distn free testing of gof

$$D1 + F1 \rightarrow v1 \rightarrow w1$$

$$D2 + F2 \rightarrow v2 \rightarrow w2$$

$$D3 + F3 \rightarrow v3 \rightarrow w3$$

$$D1 + F2 \rightarrow v4 \rightarrow w4 ??$$

KT2: invertible transforms of empirical processes between 'any'
stochastic processes

$$v1 \leftrightarrow v2 \quad v1 \leftrightarrow v3 \quad v2 \leftrightarrow v3$$

So choose v3 to be simple, generic

Is $v1 = v2$ in distn? Is v1 generated by F1?

How effective to test these after transforming into v3?

Tale of two operators

Consider the multinomially distd

$$Y_{in} = \frac{\nu_{in} - n p_i}{\sqrt{n p_i}} \quad \text{where} \quad \sum_{i=1}^m p_i = 1$$

For a suitable limiting distribution, consider $X \sim \mathcal{N}(0, I)$ and the projection orthogonal to X (Khmaladze 2013)

$$Y = (I - \sqrt{p} \sqrt{p}^T) X \quad \text{where} \quad \sqrt{p}^T = (\sqrt{p_1}, \sqrt{p_2}, \dots, \sqrt{p_m})$$

Then

$$\text{Corr}(Y_{in}) = \text{Corr}(Y) = I - \sqrt{p} \sqrt{p}^T$$

A related operator is $I - 2\sqrt{p} \sqrt{p}^T$, provided $\|\sqrt{p}\| = 1$, reflecting in an axis perpendicular to \sqrt{p} .

Reflection

To some extent, distns can be related via a common 'base' distn:

$$Y = (I - \sqrt{p}\sqrt{p}^T)X \quad Z = (I - \sqrt{r}\sqrt{r}^T)X$$

These projections involve only the second moment. But we are mainly concerned with zero mean Gaussian processes.

More generally, to move from one distn to another, define the reflection

$$U = I - \frac{2}{\|\sqrt{p} - \sqrt{r}\|^2}(\sqrt{p} - \sqrt{r})(\sqrt{p} - \sqrt{r})^T$$

whence $U\sqrt{p} = \sqrt{r}$ and $U\sqrt{r} = \sqrt{p}$.

For simple null, U maps from one empirical process to another.

For composite hypotheses with single parameter θ needing to be estimated from the data, need U to simultaneously map $\sqrt{p} \rightarrow \sqrt{r}$ and the score function $h_p \rightarrow h_r$. For $\theta = \theta_1, \theta_2, \dots, \theta_K$, need U simultaneously to map $\sqrt{p} \rightarrow \sqrt{r}$ and the K derivatives in the (gradient) score function from one distn to the other.

For sufficiently large K , the vector spaces for discrete distns may not have sufficient dimensions to accommodate all constraints on the operator U . So we group the data artificially in order to increase the available dimensionality.

A single data 'point' is then a subgroup of the original sample, which subgroup is 'rotated' into a simpler 'empirical' process.

Simulation of Bernoulli Trials

Bernoulli trials

mn independent trials, but with common distributional family of success and failure depending on a covariate X_{ij} , for

$$1 \leq i \leq m, 1 \leq j \leq n.$$

$Y_{ij} = \text{outcome} = 1$ for success, 0 for failure.

For the j th group of trials, $X_j = (X_{1j}, X_{2j}, \dots, X_{mj})$ and

$$Z_j = (Y_{1j}, Y_{2j}, \dots, Y_{mj})$$

Order the results lexicographically.

Eg: when $m = 1$,

$y = 0, 1$ corresponds to $z = 1, 2 = 2^m$.

Eg: when $m = 2$,

$(y_1, y_2) = 00, 01, 10, 11$ corresponds to $z = 1, 2, \dots, 4 = 2^m$.

Underlying joint distn given by

$$P_\theta(dx, dz) = P_{x,\theta}(dz)H(dx) \quad (1)$$

where θ is a vector of K parameters, and the distn H is unspecified. Denoting Borel sets in the respective spaces by A and B , we define the following centred empirical process:

$$\alpha_n^P(A, B, \theta) = \frac{1}{\sqrt{n}} \sum_{j=1}^n \left[\mathbb{1}_{\{Z_j \in B\}} - P_\theta(B|X_j) \right] \mathbb{1}_{\{X_j \in A\}} \quad (2)$$

Considering functions $\phi(x, z)$ in $L_2(P_\theta)$, the space of square integrable functions wrt the measure P_θ , we rewrite (2) in the form

$$\alpha_n^P(\phi, \theta) = \int \int \phi(x, z) \alpha_n^P(dx, dz, \theta) \quad (3)$$

Outline

We wish to transform processes defined on $L_2(P_\theta)$ to processes defined on say $L_2(Q_\theta)$, where by analogy to (1) we have

$$Q_\theta(dx, dz) = Q_{x,\theta}(dz)H(dx)$$

Consider functions $\phi \in L_2(P_\theta)$ and $\psi \in L_2(Q_\theta)$. Provided that P_θ and Q_θ are equivalent, define

$$\ell_{x,\theta}(dz) = \sqrt{\frac{dQ_{x,\theta}(dz)}{dP_{x,\theta}(dz)}} \quad (4)$$

so that $\psi \in L_2(Q_\theta)$ if and only if $\ell\psi \in L_2(P_\theta)$.

We simplify the conditional $Q_{x,\theta}(y)$ measure so that the numerator of the quantity under the square root in (4) depends on neither x nor θ : we may then write

$$\ell_{x,\theta}(z) = \sqrt{\frac{\prod_{i=1}^m q^{1-y_i} (1-q)^{y_i}}{\prod_{i=1}^m p_{x,\theta}^{1-y_i} (1-p_{x,\theta})^{y_i}}} \quad (5)$$

with q and $p_{x,\theta}$ denoting probabilities of failure, and $y = 1$ indicating success.

The transformation of the empirical process on $L_2(P_{x,\theta})$ to that on $L_2(Q)$ is effected as follows. From Khmaladze (2017) and (3),

$$\alpha_n^Q(\psi, \theta) = \alpha_n^Q(\psi) = \alpha_n^P(U\ell\psi, \theta) \quad (6)$$

in which U is a unitary operator.

We wish to carry out goodness of fit tests as to whether or not the data in (2) could reasonably have been generated by P_θ , for some value of θ . A standard means of doing this would be to employ the KS statistic, defined from (3) as

$$\max_{\phi \in \Phi} |\alpha_n^P(\phi, \theta)| \quad (7)$$

where the maximum is taken over all functions ϕ in a class Φ of functions in $L_2(P_\theta)$.

Kolmogorov Smirnov test on rotated distn

The point of transforming to the Q space, however, is that we may instead employ the KS statistic in the Q space:

$$\max_{\psi \in \Psi} |\alpha_n^Q(\psi)| \quad (8)$$

and test goodness of fit by an empirical process defined on a simpler measure Q depending on neither x nor θ .

A suitable class Ψ is generated by indicator functions in x and z :

$$\psi_{x_0, z_0}(x, z) = \mathbb{1}_{\{x \leq x_0\}} \left[\mathbb{1}_{\{z \leq z_0\}} - 2^{-m} z_0 \right] \quad (9)$$

in which ψ is regarded as a vector of length 2^m , so that the operator U assumes the form of a $2^m \times 2^m$ orthogonal matrix.

Target distn

The original process is generated by P , with normalised score functions a_1, a_2, \dots, a_K : the vectors $1, a_1, a_2, \dots, a_K$ are mutually orthogonal (wrt the P measure). The target process has $1, b_1, b_2, \dots, b_K$ mutually orthogonal wrt the Q measure. As the target process for the Bernoulli trials we choose the logistic distribution with parameter set to zero, and covariates to unity. For $m = 1$, the target score function becomes $b_1 = (-1, 1)$; while for $m = 2, 3$ and 4 , b_1 becomes respectively

$$\sqrt{2}(-1, 0, 0, 1) \quad \frac{1}{\sqrt{3}}(-3, -1, -1, 1, -1, 1, 1, 3)$$

$$\frac{1}{\sqrt{4}}(-4, -2, -2, 0; -2, 0, 0, 2; -2, 0, 0, 2; 0, 2, 2, 4)$$

More generally, however, the target score functions can be arbitrary, save for the orthogonality constraints?

Choosing covariates

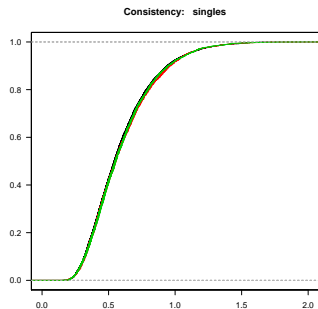
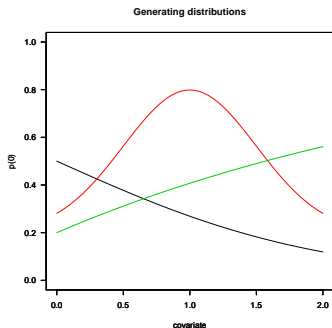


Figure:

Choosing covariates for Bernoulli trials

consistency: known θ , 96 trials, singles, 5000 replications

Consistency

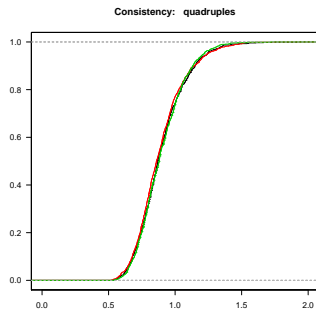
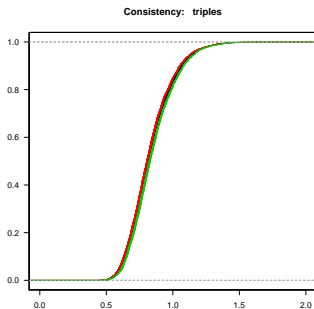


Figure:

consistency: unknown θ , 96 trials, triples, 5000 replications

consistency: unknown θ , 96 trials, quadruples, 1000 replications

Power

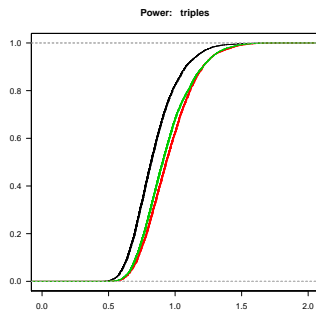
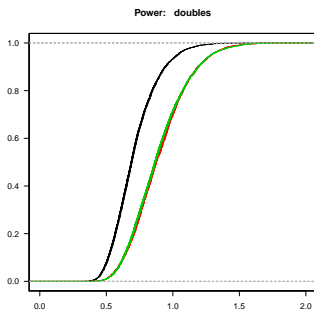


Figure:

Power, unknown θ , 96 trials, doubles, 5000 replications

Power, unknown θ , 96 trials, triples, 5000 replications

Conclusion

advantages of KT2

- intuitive
 - for discrete distns, linear algebra suffices for basic intuition
- easy to code, computing fast

So why has KT2 been so little applied?

Nguyen (2017*a,b*) successfully applies KT2 to contingency tables and distributional tails; but she does not have covariates changing over the sample, and nor does she allow for estimating unknown parameters.

For this simple simulation of Bernoulli trials

- We have verified consistent mappings of distinct empirical processes into common base distn
- and shown high power to reject incorrect null hypotheses

- $KT = KT1$ beautifully elegant way to test gof by distn free methods, but technically difficult and hard to implement
- $KT2$ reflections or 'rotations' easily implemented
- Testing after 'rotation' of empirical processes to that of a simple process seems effective
 - possible need to artificially group data for rotation
 - but this seems not to make the $KT2$ 'rotation into distn free' methods less effective.
- This paper is apparently the first to apply $KT2$ when distns depend on exogenous covariates, and the first to investigate empirically the working of the methodology when parameters need to be estimated.

- All sorts of outstanding issues
 - tie in with other mathematical techniques? eg, reflections \leftrightarrow root systems of Lie algebras?
 - mapping between distns with varying number of parameters?
 - Are the score functions of target really so arbitrary?
 - Testing constraints on parameters after rotation?
 - Using other gof tests, eg ω^2 , as well as KS statistics after rotation.
 - Large number of parameters?
 - KT2 good for this simple simulation of Bernoulli trials. For more general situations?
 - Clarify minimum group size for given modelling situations and number of parameters. Optimal group size?
Preferable to have X well dispersed within groups?
- WHAT are we to call KT2?

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- Roberts, L. A. (2016), On distribution free testing in Bernoulli trials with distributions depending on covariates.
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Thank you

and also thanks to Estate Khmaladze

as well as apologies to Charles Dickens